

Approximate Bayesian computation: likelihood-free inference for complex models

Richard Wilkinson

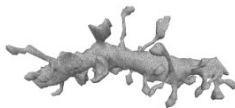
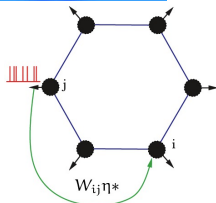
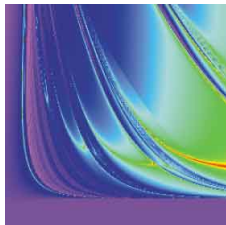
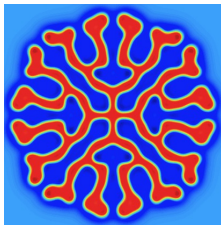
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Calibration

- For most simulators we specify parameters θ and i.c.s and the simulator, $f(\theta)$, generates output X .
- The inverse-problem: observe data D , estimate parameter values θ which explain the data.

The inverse/ calibration/ parameter estimation/... problem is estimating θ that could have led to D



Statistical inference

Consider the following three parts of inference:

1 Modelling

2 Inferential framework

3 Statistical computation

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- ▶ Simulator - generative model $\pi(X|\theta)$
- ▶ Statistical model
 - ★ prior distributions on unknown parameters, $\pi(\theta)$
 - ★ observation error on the data, $\pi(D|X)$
 - ★ simulator error (if its not a perfect representation of reality)

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- ▶ Classical/frequentist
- ▶ Bayesian
- ▶ History matching

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- ▶ this remains hard even with increased computational resource

Inferential framework

Classical/frequentist

- Maximum likelihood

$$\hat{\theta} = \arg \max_{\theta} \pi(D|\theta)$$

or a more ad-hoc approach

$$\hat{\theta} = \arg \min_{\theta} (\mathbb{E}(D|\theta) - D)^2$$

- Can find confidence intervals (with coverage guarantees etc)
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Bayesian

- Work only with probabilities (no significance, confidence, p-values)
- update beliefs in light of data and aim to find posterior distributions

$$\pi(\theta|D) \propto \pi(\theta)\pi(D|\theta)$$

posterior \propto prior \times likelihood

- Needs a prior distribution, computation is still hard but often do able

Computational Intractability

$$\pi(\theta|D) = \frac{\pi(D|\theta)\pi(\theta)}{\pi(D)}$$

- **usual intractability** in Bayesian inference is not knowing $\pi(D)$.
- a problem is **doubly intractable** if $\pi(D|\theta) = c_\theta p(D|\theta)$ with c_θ unknown
- a problem is **completely intractable** if $\pi(D|\theta)$ is unknown and can't be evaluated (unknown is subjective). I.e., if the analytic distribution of the simulator, $f(\theta)$, run at θ is unknown.

Completely intractable models are where we need to resort to ABC methods

Approximate Bayesian Computation (ABC)

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- they do not require explicit knowledge of the likelihood function
- inference is done using simulation from the model (they are 'likelihood-free').

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ABC methods are popular in biological disciplines as they are

- Simple to implement
- Intuitive
- Embarrassingly parallelizable
- Can usually be applied

Rejection ABC

Uniform Rejection Algorithm

- Draw θ from $\pi(\theta)$
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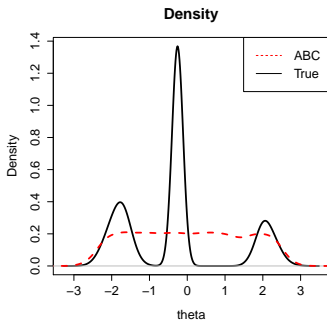
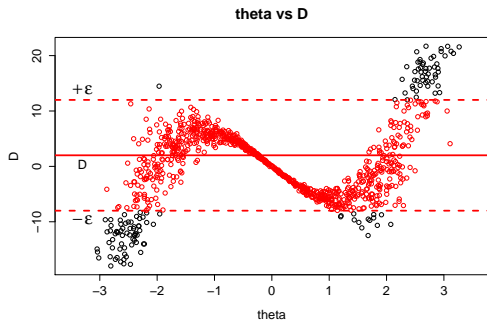
ϵ reflects the tension between computability and accuracy.

- As $\epsilon \rightarrow \infty$, we get observations from the prior, $\pi(\theta)$.
- If $\epsilon = 0$, we generate observations from $\pi(\theta | D)$.

Rejection sampling is inefficient, but we can adapt other MC samplers such as MCMC and SMC.

Simple \rightarrow Popular with non-statisticians

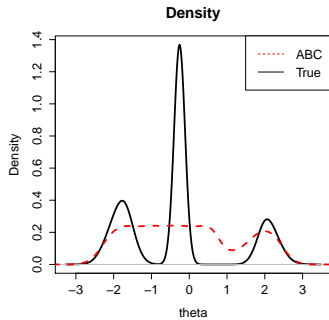
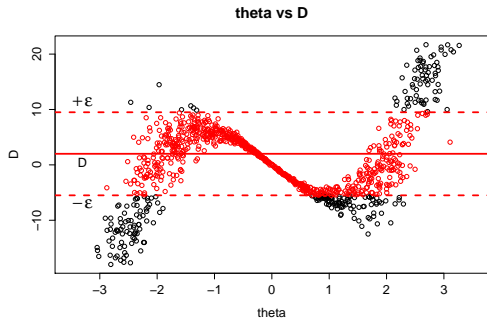
$$\epsilon = 10$$



$$\theta \sim U[-10, 10], \quad X \sim N(2(\theta + 2)\theta(\theta - 2), 0.1 + \theta^2)$$

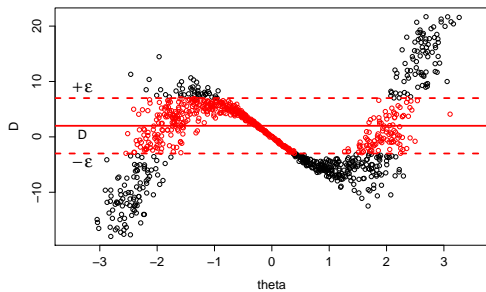
$$\rho(D, X) = |D - X|, \quad D = 2$$

$$\epsilon = 7.5$$

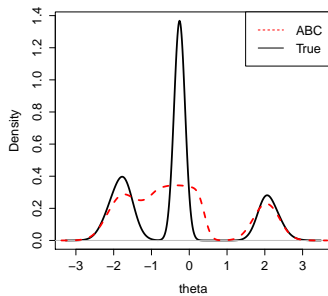


$$\epsilon = 5$$

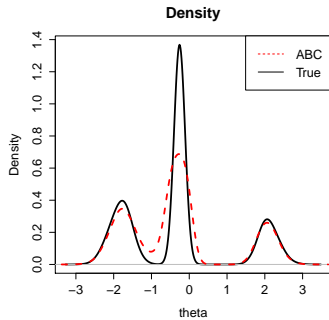
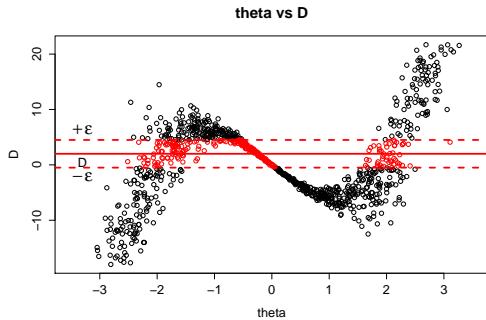
theta vs D



Density

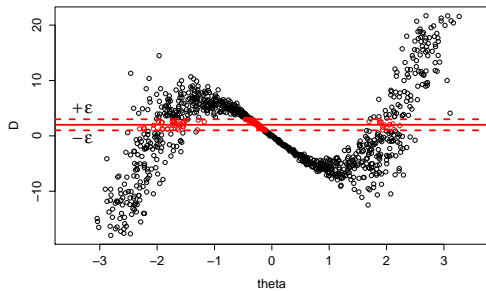


$$\epsilon = 2.5$$

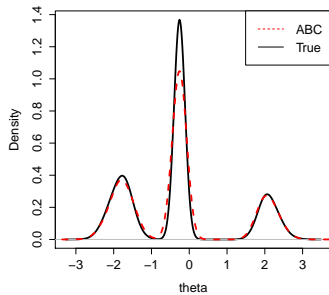


$$\epsilon = 1$$

theta vs D



Density



Rejection ABC

If the data are too high dimensional we never observe simulations that are 'close' to the field data - **curse of dimensionality**

Reduce the dimension using summary statistics, $S(D)$.

Approximate Rejection Algorithm With Summaries

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- Summary statistic $S(D)$ - controls 'information loss'
 - ▶ inference is based on $\pi(\theta|S(D))$ rather than $\pi(\theta|D)$
 - ▶ a combination of expert judgement, and stats/ML tools can be used to find informative summaries

Computation

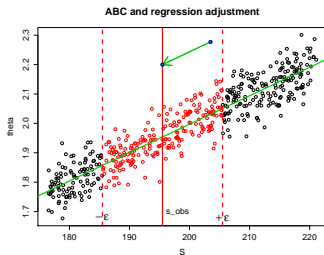
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- Post-hoc corrections



use the estimate of the posterior mean at s_{obs} and the residuals from the fitted line to form the posterior.

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The first approximation allows the matching between $S(D)$ and $S(X)$ to be done in a lower dimension. There is a trade-off

- $\dim(S)$ small: $\pi(\theta|S_{obs}) \approx \pi_{ABC}(\theta|S_{obs})$, but $\pi(\theta|S_{obs}) \not\approx \pi(\theta|D)$
- $\dim(S)$ large: $\pi(\theta|S_{obs}) \approx \pi(\theta|D)$ but $\pi(\theta|S_{obs}) \not\approx \pi_{ABC}(\theta|S_{obs})$ as curse of dimensionality forces us to use larger ϵ

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Optimal (in some sense) to choose $\dim(s) = \dim(\theta)$

Choosing summary statistics

If $S(D) = s_{obs}$ is sufficient for θ , i.e., s_{obs} contains all the information contained in D about θ

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- Recent progress made with random forest and neural-network models to learn the relevant features
 - 1 Train a ML model, $m(D)$, to predict θ from D using a large number of simulator runs $\{\theta_i, D_i\}$
 - 2 ABC then simulates θ from the prior and D from the simulator, and accepts θ if $m(D) \approx m(D_{obs})$

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- Ignoring discrepancy can lead to over-confident and incorrect inference about θ
- When using ABC, you are automatically including some characterization of model discrepancy (determined by the summaries, metric and tolerance you chose).
 - ▶ So it's better to have thought carefully about this.
 - ▶ May only be a case of thinking about an approximate magnitude of the discrepancy

History matching

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where $\text{Var}(D|\theta)$ is the total variance taking into account measurement error, discrepancy, emulator uncertainty etc.

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HM is a conservative approach - it only rules out parameters we are reasonably confident are improbable. It doesn't attempt to tell us the best parameter value.

Conclusions

ABC allows inference in models for which it would otherwise be impossible.

- not a silver bullet - if likelihood methods possible, use them instead.

Efficient algorithms and post-hoc regression adjustments can greatly improve computational efficiency, but computation is still usually the limiting factor.

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Thank you for listening!